Machine Learning and Statistics in Genetics and Genomics VI: Introduction to Gaussian Processes

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Microsoft Research eScience group Research

Los Angeles , USA

Current topics in computational biology UCLA Winter quarter 2014

Motivation

Intuitive approach

Function space view

Outline

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Function space view

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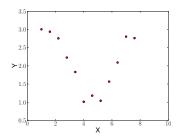
Why Gaussian processes?

- So far: linear models with a finite number of basis functions, e.g. φ(x) = (1, x, x²,...,x^K)
- Open questions:
 - How to design a suitable basis?
 - How many basis functions to pick?

 Gaussian processes: accurate and flexible regression method yielding predictions alongside with error bars.

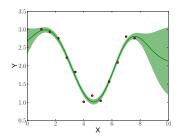
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Linear model, accounting for a set of measured SNPs X

$$p(\boldsymbol{y} \mid \boldsymbol{X}, \boldsymbol{\theta}, \sigma^2) = \mathcal{N}\left(\boldsymbol{y} \mid \sum_{s=1}^{S} \boldsymbol{x}_s \theta_s, \sigma^2 \boldsymbol{I}\right)$$

► Prediction at unseen test input given max. likelihood weight: $p(y^* | \boldsymbol{x}^*, \hat{\boldsymbol{\theta}}) = \mathcal{N}\left(y^* | \boldsymbol{x}^* \hat{\boldsymbol{\theta}}, \sigma^2\right)$

Marginal likelihood

$$p(\boldsymbol{y} \mid \boldsymbol{X}, \sigma^{2}, \sigma_{g}^{2}) = \int_{\boldsymbol{\theta}} \mathcal{N}\left(\boldsymbol{y} \mid \boldsymbol{X}\boldsymbol{\theta}, \sigma^{2}\boldsymbol{I}\right) \mathcal{N}\left(\boldsymbol{\theta} \mid \boldsymbol{0}, \sigma_{g}^{2}\boldsymbol{I}\right)$$
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Further reading

- C. E. Rasmussen, C. K. Williams Gaussian processes for machine learning
 - Comprehensive and freely available introduction (Appendix!).
- ► Christopher M. Bishop: Pattern Recognition and Machine learning

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The Gaussian distribution

Gaussian processes are merely based on the good old Gaussian

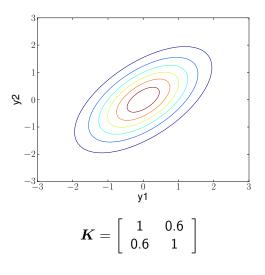
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Covariance matrix or kernel matrix

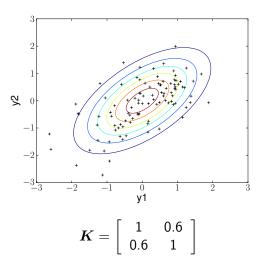
Probability contour

Samples



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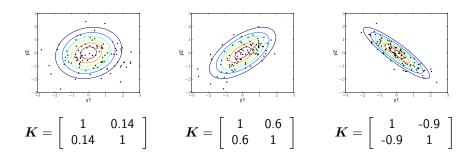
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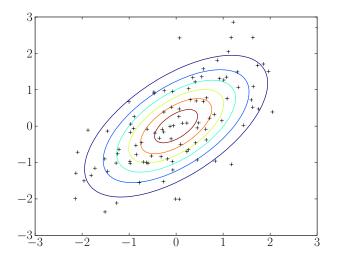
Varying the covariance matrix



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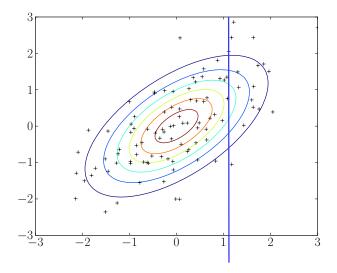
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Inference



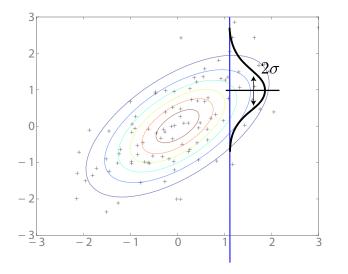
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Inference

- ▶ Joint probability $p(y_1, y_2 \mid \mathbf{K}) = \mathcal{N}([y_1, y_2] \mid \mathbf{0}, \mathbf{K})$
- Conditional probability

$$p(y_2 \mid y_1, \mathbf{K}) = \frac{p(y_1, y_2 \mid \mathbf{K})}{p(y_1 \mid \mathbf{K})}$$
$$\propto \exp\left\{-\frac{1}{2}[y_1, y_2] \mathbf{K}^{-1} \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}\right\}$$

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 Completing the square yields a Gaussian with non-zero as posterior for y₂.

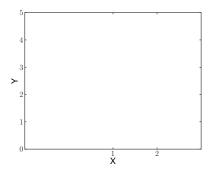
Inference

Gaussian conditioning in 2D

$$\begin{split} p(y_2 \mid y_1, \mathbf{K}) &= \frac{p(y_1, y_2 \mid \mathbf{K})}{p(y_1 \mid \mathbf{K})} \propto \exp\left\{-\frac{1}{2}[y_1, y_2] \, \mathbf{K}^{-1} \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}\right\} \\ &= \exp\{-\frac{1}{2} \begin{bmatrix} y_1^2 \mathbf{K}_{1,1}^{-1} + y_2^2 \mathbf{K}_{2,2}^{-1} + 2y_1 \mathbf{K}_{1,2}^{-1} y_2 \end{bmatrix}\} \\ &= \exp\{-\frac{1}{2} \begin{bmatrix} y_2^2 \mathbf{K}_{2,2}^{-1} + 2y_2 \mathbf{K}_{1,2}^{-1} y_1 + C \end{bmatrix}\} \\ &= Z \exp\{-\frac{1}{2} \mathbf{K}_{2,2}^{-1} \begin{bmatrix} y_2^2 + 2y_2 \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}} \end{bmatrix}\} \\ &= Z \exp\{-\frac{1}{2} \mathbf{K}_{2,2}^{-1} \begin{bmatrix} y_2^2 + 2y_2 \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}} \end{bmatrix}\} \\ &= Z \exp\{-\frac{1}{2} \mathbf{K}_{2,2}^{-1} \begin{bmatrix} y_2^2 + 2y_2 \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}} \end{bmatrix}\} \\ &= Z \exp\{-\frac{1}{2} \mathbf{K}_{2,2}^{-1} \begin{bmatrix} y_2 + 2y_2 \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}} \end{bmatrix} + \frac{1}{2} \mathbf{K}_{2,2}^{-1} \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}}^2 \} \\ &= Z' \exp\{-\frac{1}{2} \underbrace{\mathbf{K}_{2,2}^{-1}}_{\sigma^2} \begin{bmatrix} y_2 + \frac{\mathbf{K}_{1,2}^{-1} y_1}{\mathbf{K}_{2,2}^{-1}} \end{bmatrix}^2\} \propto \mathcal{N}\left(y_2 \mid \mu, \sigma^2\right) \end{split}$$

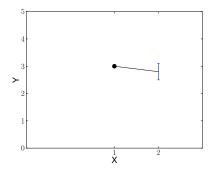
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- Let us interpret y_1 and y_2 as outputs in a regression setting.
- We can introduce an additional 3rd point.



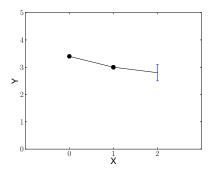
Now P([y₁, y₂, y₃] | K₃) = N ([y₁, y₂, y₃] | 0, K₃), where K₃ is now a 3 x 3 covariance matrix!

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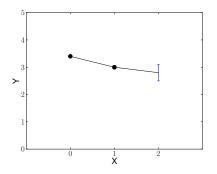
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Now $P([y_1, y_2, y_3] | \mathbf{K}_3) = \mathcal{N}([y_1, y_2, y_3] | \mathbf{0}, \mathbf{K}_3)$, where \mathbf{K}_3 is now a 3 x 3 covariance matrix!

- Analogously we can look at the joint probability for arbitrary many points and obtain predictions.
- Issue: how to construct a good covariance matrix?

• A simple heuristics

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- The ordering of the points y_1, y_2, y_3 matters.
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A general recipe

▶ Use a covariance function (kernel function) to construct *K*:

$$\boldsymbol{K}_{i,j} = k(x_i, x_j; \boldsymbol{\Theta}_{\mathsf{K}})$$

Example: The linear covariance function corresponds to a variance component model

$$k_{\mathsf{LIN}}(x_i, x_j, ; A) = A^2 x_i \cdot x_j$$

Example: The squared exponential covariance function embodies the belief that points further apart are less correlated:

$$k_{\text{SE}}(x_i, x_j; A, L) = A^2 \exp\left\{-0.5 \cdot \frac{(x_i - x_j)^2}{L^2}\right\}$$

• $\boldsymbol{\Theta}_{\mathsf{K}} = \{A, L\}$: hyperparameters.

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 ► Denote the covariance matrix for a set of inputs X = {x₁,...,x_N} as: K_{X,X}(Θ_K)

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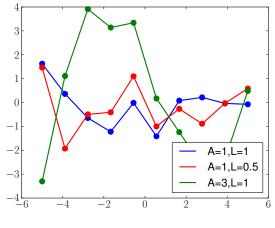
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GP samples using the squared exponential covariance function

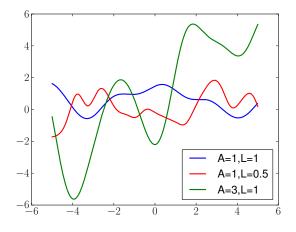


10D Gaussian

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Constructing Covariance Matrices

GP samples using the squared exponential covariance function

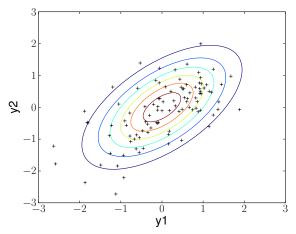


500D Gaussian

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Constructing Covariance Matrices

GP samples using the squared exponential covariance function



Reminder: Every function line corresponds to a sample drawn from this 2D Gaussian!

Drawing samples from a Gaussian processes

For each sample do:

- Choose discretization of x axes $X = \{x_0, x_1, \dots, x_N\}$.
- Evaluate covariance $\boldsymbol{K} = \boldsymbol{K}_{\boldsymbol{X},\boldsymbol{X}}(\boldsymbol{\varTheta}_{\mathsf{K}})$



"Matlab"

 Draw independent Gaussian variables

 $\tilde{\pmb{y}} = \mathsf{randn}(N,1)$

• Rotate with \sqrt{K}

 $oldsymbol{y} = \mathsf{chol}(oldsymbol{K}) \cdot \widetilde{oldsymbol{y}}$

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Math

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- Consistency of the 10D and 500D Gaussian.
- A small quiz:
 - Let y_1, y_2, y_3 have covariance matrix

$$\boldsymbol{K}_3 = \left[\begin{array}{cccc} 1 & 0.5 & 0 \\ 0.5 & 1 & 0.5 \\ 0 & 0.5 & 1 \end{array} \right] \text{ and inverse } \boldsymbol{K}_3^{-1} = \left[\begin{array}{cccc} 1.5 & -1 & 0.5 \\ -1 & 2 & -1 \\ 0.5 & -1 & 1.5 \end{array} \right]$$

i.e.
$$p(\{y_1, y_2, y_3\} | \mathbf{K}_3) = \mathcal{N}(\{y_1, y_2, y_3\} | \mathbf{0}, \mathbf{K}_3)$$

Now focus on the variables y_1, y_2 , integrating out y_3

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Which of the following statements is true

a)
$$K_2 = \begin{bmatrix} 1 & 5 \\ 5 & 1 \end{bmatrix}$$
 b) $K_2^{-1} = \begin{bmatrix} 1.5 & -1 \\ -1 & 2 \end{bmatrix}$

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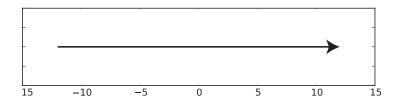
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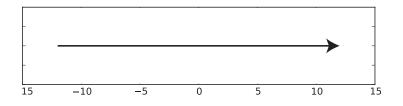
GP as infinite object (philosophical)

- ► A valid covariance function k(x, x') defines recipe to calculate covariance for any choice of inputs.
- Prior on functions: all points on the real line are inputs; K_{R,R} is an infinite object!
- Numerical implementation: choose finite subset X and evaluate on a reduced, finite K_{X,X}, exploiting consistency rule.



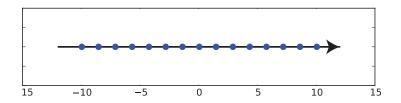
GP as infinite object (philosophical)

- ► A valid covariance function k(x, x') defines recipe to calculate covariance for any choice of inputs.
- ► Prior on functions: all points on the real line are inputs; K_{R,R} is an infinite object!
- ▶ Numerical implementation: choose finite subset *X* and evaluate on a reduced, finite *K*_{*X*,*X*}, exploiting consistency rule.



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Outline

Motivation

Intuitive approach

Function space view

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Function space view

So far

- 1. Joint Gaussian distribution over the set of all outputs y.
- 2. Covariance function as a recipe to construct a suitable covariance matrices from the corresponding inputs X.

Function space view

The Gaussian process as a prior on functions

- Covariance function and hyperparameters reflect the prior belief on function smoothness, lengthscales etc.
- The general recipe allows a joint Gaussian to be constructed for an arbitrary selection of input locations X.

Prior on infinite function f(x)

 $p(f(x)) = \mathsf{GP}(f(x) \,|\, k)$

Prior on function values $f = (f_1, \dots, f_N)$

 $p(\boldsymbol{f} \mid \boldsymbol{X}, \boldsymbol{\Theta}_{\mathsf{K}}) = \mathcal{N}\left(\left. \boldsymbol{f} \mid \boldsymbol{0}, \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}} \right| \boldsymbol{\Theta}_{\mathsf{K}}
ight)$

Noise-free observations

- Given noise-free training data $\mathcal{D} = \{ \boldsymbol{x}_n, f_n \}_{n=1}^N$
- \blacktriangleright Want to make predictions f^{\star} at test points X^{\star}
- \blacktriangleright Joint distribution of f and f^{\star} is

$$p([\boldsymbol{f}, \boldsymbol{f}^{\star}] \,|\, \boldsymbol{X}, \boldsymbol{X}^{\star}, \boldsymbol{\Theta}_{\mathsf{K}}) = \mathcal{N}\left(\begin{bmatrix} \boldsymbol{f}, \boldsymbol{f}^{\star} \end{bmatrix} \,|\, \boldsymbol{0}, \left[\begin{array}{cc} \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}^{\star}} \\ \boldsymbol{K}_{\boldsymbol{X}^{\star}, \boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}^{\star}, \boldsymbol{X}^{\star}} \end{array} \right] \right)$$

(All kernel matrices K depend on hyperparameters Θ_{K} which are dropped for brevity.)

Real data is rarely noise-free.

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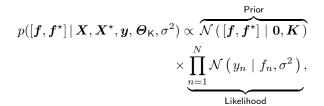
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▶ Given observed noisy data $D = \{X, y\}$, the joint probability over latent function values f and f^* given y is

$$p([\boldsymbol{f}, \boldsymbol{f}^{\star}] | \boldsymbol{X}, \boldsymbol{X}^{\star}, \boldsymbol{y}, \boldsymbol{\Theta}_{\mathsf{K}}, \sigma^{2}) \propto \mathcal{N}\left([\boldsymbol{f}, \boldsymbol{f}^{\star}] \mid \boldsymbol{0}, \begin{bmatrix} \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}^{\star}} \\ \boldsymbol{K}_{\boldsymbol{X}^{\star}, \boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}^{\star}, \boldsymbol{X}^{\star}} \end{bmatrix}\right) \times \underbrace{\prod_{n=1}^{N} \mathcal{N}\left(y_{n} \mid f_{n}, \sigma^{2}\right)}_{\mathsf{Likelihood}},$$

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 \blacktriangleright Applying "Gaussian calculus", integrating out f yields

$$p([\boldsymbol{y},\boldsymbol{f}^{\star}] \mid \boldsymbol{X}, \boldsymbol{X}^{\star}, \boldsymbol{y}, \boldsymbol{\Theta}_{\mathsf{K}}, \sigma^{2}) \propto \mathcal{N}\left(\begin{bmatrix}\boldsymbol{y},\boldsymbol{f}^{\star}\end{bmatrix} \mid \boldsymbol{0}, \begin{bmatrix} \boldsymbol{K}_{\boldsymbol{X},\boldsymbol{X}} + \sigma^{2}\boldsymbol{I} & \boldsymbol{K}_{\boldsymbol{X},\boldsymbol{X}^{\star}} \\ \boldsymbol{K}_{\boldsymbol{X}^{\star},\boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}^{\star},\boldsymbol{X}^{\star}} \end{bmatrix}\right)$$

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Note: Assuming noisy instead of perfect observation noise merely corresponds to adding a diagonal component to the self-covariance K_{X,X}.

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Making predictions

 The predictive distribution follows from the joint distribution by completing the square (conditioning)

$$p([\boldsymbol{y},\boldsymbol{f}^{\star}] \mid \boldsymbol{X}, \boldsymbol{X}^{\star}, \boldsymbol{y}, \boldsymbol{\Theta}_{\mathsf{K}}, \sigma^{2}) \propto \mathcal{N}\left([\boldsymbol{y},\boldsymbol{f}^{\star}] \mid \boldsymbol{0}, \begin{bmatrix} \boldsymbol{K}_{\boldsymbol{X},\boldsymbol{X}} + \sigma^{2}\boldsymbol{I} & \boldsymbol{K}_{\boldsymbol{X},\boldsymbol{X}^{\star}} \\ \boldsymbol{K}_{\boldsymbol{X}^{\star},\boldsymbol{X}} & \boldsymbol{K}_{\boldsymbol{X}^{\star},\boldsymbol{X}^{\star}} \end{bmatrix}\right)$$

• Gaussian predictive distribution for f^* $p(f^* | X, y, X^*, \Theta_K, \sigma^2) = \mathcal{N} (f^* | \mu^*, \Sigma^*)$ with $\mu^* = K_{X^*, X} [K_{X, X} + \sigma^2 I]^{-1} y$ $\Sigma^* = K_{X^*, X^*} - K_{X^*, X} [K_{X, X} + \sigma^2 I]^{-1} K_{X, X^*}$

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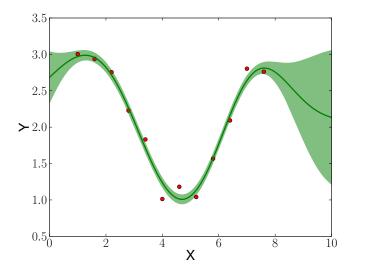
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• Gaussian predictive distribution for f^{\star}

$$p(f^{\star} | \boldsymbol{X}, \boldsymbol{y}, \boldsymbol{X}^{\star}, \boldsymbol{\Theta}_{\mathsf{K}}, \sigma^{2}) = \mathcal{N} (f^{\star} | \boldsymbol{\mu}^{\star}, \boldsymbol{\Sigma}^{\star})$$
 with
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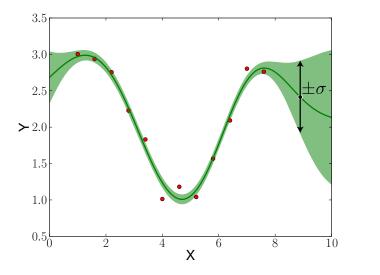
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Making predictions Example



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Making predictions Example



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- 1. Fixed covariance matrix: $p(\boldsymbol{y} \,|\, \boldsymbol{K})$
- 2. Constructed covariance matrix: $\{K\}_{i,j} = k(x_i, x_j; \boldsymbol{\Theta}_{\mathsf{K}})$

3. Can we learn the hyperparameters Θ_{K} ?

Formally we are interested in the posterior

 $p(\boldsymbol{\Theta}_{\mathsf{K}} \,|\, \mathcal{D}) \propto p\left(\boldsymbol{y} \,|\, \boldsymbol{X}, \, \boldsymbol{\Theta}_{\mathsf{K}}\right) p(\boldsymbol{\Theta}_{\mathsf{K}})$

- Inference is analytically intractable!
- MAP estimate instead of a full posterior. Set \(\mathcal{O}_K\) to the most probable hyperparameter settings:

$$\begin{split} \hat{\boldsymbol{\vartheta}}_{\mathsf{K}} &= \operatorname*{argmax}_{\boldsymbol{\Theta}_{\mathsf{K}}} \ln \left[p\left(\boldsymbol{y} \mid \boldsymbol{X}, \boldsymbol{\Theta}_{\mathsf{K}} \right) p(\boldsymbol{\Theta}_{\mathsf{K}}) \right] \\ &= \operatorname*{argmax}_{\boldsymbol{\Theta}_{\mathsf{K}}} \ln \mathcal{N} \left(\boldsymbol{y} \mid \boldsymbol{0}, \boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}}(\boldsymbol{\Theta}_{\mathsf{K}}) + \sigma^{2} \boldsymbol{I} \right) + \ln p(\boldsymbol{\Theta}_{\mathsf{K}}) \\ &= \operatorname*{argmax}_{\boldsymbol{\Theta}_{\mathsf{K}}} \left[-\frac{1}{2} \log \det[\boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}}(\boldsymbol{\Theta}_{\mathsf{K}}) + \sigma^{2} \boldsymbol{I}] \\ &- \frac{1}{2} \boldsymbol{y}^{\top} [\boldsymbol{K}_{\boldsymbol{X}, \boldsymbol{X}}(\boldsymbol{\Theta}_{\mathsf{K}}) + \sigma^{2} \boldsymbol{I}]^{-1} \boldsymbol{y} - \frac{N}{2} \log 2\pi + \ln p(\boldsymbol{\Theta}_{\mathsf{K}}) \right] \end{split}$$

 Optimization can be carried out using standard optimization techniques.

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 Optimization can be carried out using standard optimization techniques.

- ► The covariance function embodies the prior belief about functions.
- Example: linear regression

$$y_n = wx_n + c + \psi_n$$

Covariance function denote covariation

$$k(x_n, x'_n) = \langle y_n y'_n \rangle$$

= $\langle (wx_n + c + \psi_n)(wx'_n + c + \psi'_n) \rangle$
= $\underbrace{w^2 \cdot x_n x'_n + c^2}_{\text{kernel: } k(x_n, x'_n)} + \delta_{n,n'} \psi_n^2$

Multidimensional input space

- Generalise squared exponential covariance function to multiple dimensions
 - ► 1 Dimension $k_{\text{SE}}(x_i, x_j; A, L) = A^2 \exp\left\{-0.5 \cdot \frac{(x_i x_j)^2}{I^2}\right\}$
 - ► D Dimensions dD $k_{\mathsf{SE}}(x_i, x_j; A, L) = A^2 \exp\left\{-0.5 \sum_{d=1}^{D} \frac{(x_i^d - x_j^d)^2}{L_d^2}\right\}$
- Lengthscale parameters L_d denote "relevance" of a particular data dimension.

• Large L_d correspond to irrelevant dimensions.

Multidimensional input space

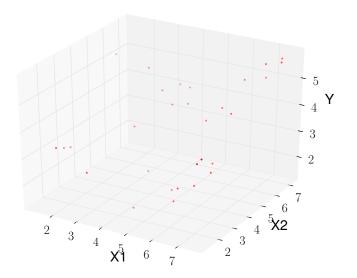
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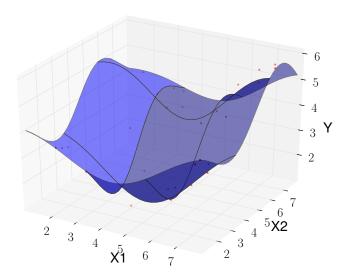
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2D regression



2D regression



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Any kernel will do

Established kernels are all valid covariance functions, allowing for a wide range of possible input domains X:

- Graph kernels (molecules)
- Kernels defined on strings (DNA sequences)

Combining existing covariance functions

The sum of two covariances functions is itself a valid covariance function

$$k_S(x, x') = k_1(x, x') + k_2(x, x')$$

The product of two covariance functions is itself a valid covariance function

$$k_P(x, x') = k_1(x, x') \cdot k_2(x, x')$$

Variance component

Linear model

$$p(\boldsymbol{y} \mid \boldsymbol{X}, \boldsymbol{\theta}, \sigma^2) \\ = \mathcal{N} \left(\boldsymbol{y} \mid \boldsymbol{\Phi}(\boldsymbol{X}) \cdot \boldsymbol{\theta}, \sigma^2 \boldsymbol{I} \right)$$

• Marginalize over θ

 $p(\boldsymbol{y} \mid \boldsymbol{X}, \sigma_g^2, \sigma^2) = \mathcal{N}(\boldsymbol{y} \mid \boldsymbol{0}, \underbrace{\sigma_g^2 \boldsymbol{\Phi}(\boldsymbol{X}) \boldsymbol{\Phi}(\boldsymbol{X})^{\mathsf{T}}}_{K} + \sigma^2 \boldsymbol{I})$

Gaussian process

- ▶ Define covariance through "recipe" K_{X,X}(Θ_K)
- Implies marginal likelihood

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The inverse is not necessarily true!

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